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Altruism, fertility, and the value of children: Health policy evaluation and intergenerational welfare [☆]

Javier A. Birchenall ^{a,*}, Rodrigo R. Soares ^b

^a University of California at Santa Barbara, United States

^b Catholic University of Rio de Janeiro, University of Maryland, NBER, and IZA, Brazil

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ABSTRACT

This paper accounts for the value of children and future generations in the evaluation of health policies. This is achieved through the incorporation of altruism and fertility in a “value of life” type of framework. We are able to express adults' willingness to pay for changes in child mortality and also to incorporate the welfare of future generations in the evaluation of current policies. Our model clarifies a series of puzzles from the literature on the “value of life” and on intergenerational welfare comparisons. We show that, by incorporating altruism and fertility into the analysis of the recent U.S. experience, the estimated welfare gain of a young adult from reductions in mortality easily doubles.

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1. Introduction

This paper accounts for the value of children and future generations in the evaluation of health policies. This is achieved through the incorporation of altruism and fertility in a “value of life” type of framework. With altruism and fertility, the value that parents attach to children becomes intrinsically related to the value that present generations attach to the welfare of future generations. By bringing these new dimensions into the analysis, we are able to express adults' willingness to pay for changes in child mortality and to incorporate the welfare of future generations in the evaluation of current policies. This is done in a setting where fertility, and therefore the size of future generations itself, is endogenous to current generations' decisions. Our model clarifies a series of puzzles from the literature on the “value of life” and on intergenerational welfare comparisons, ranging from the profile of the “value of a statistical life” over the life-cycle to the specific way in which future generations—whose existence itself may depend on current actions—could be incorporated in cost–benefit analyses.

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* Corresponding author.

E-mail addresses: jbirche@econ.ucsb.edu (J.A. Birchenall), soares@econ.umd.edu (R.R. Soares).

An increasing body of literature has applied the “value of life” methodology to analyze different aspects of welfare gains from changes in mortality rates. These studies made important contributions to the understanding and measurement of non-monetary aspects of human welfare. Yet they suffer from an important methodological limitation: they do not incorporate in the analysis altruism, fertility decisions, and their consequences in terms of welfare evaluation.¹ Two interrelated dimensions of fertility choice are potentially important in evaluating life expectancy and health-related welfare gains. First, child mortality rates can be very important in determining welfare in a context in which individuals choose the number of children they have. Second, under altruistic preferences, life expectancy gains have a twofold effect: they directly increase utility via increased survival probabilities, and they increase utility via increased welfare of the offspring.

In this paper, we evaluate the welfare implications of mortality reductions in a setup in which individuals choose the number of children they have and are altruistic toward their children. We show that, under these circumstances, the value of adult mortality changes can be decomposed into three factors: the consumption factor from the traditional “value of life” specification, discussed in Rosen (1988); a fertility factor, which accounts for the welfare improvements related to the higher probability of having children; and an altruism factor, which accounts for the fact that mortality reductions will also be enjoyed by all future generations. In addition, our approach allows us to calculate an adult’s willingness to pay for reductions in child mortality. This willingness to pay depends on the effect of child mortality on the final costs of child production, the uncertainty regarding the number of surviving children, and the emotional loss associated with the death of a child. As in the case of adult mortality, child mortality reductions are also valued because future generations will benefit from them. In fact, we show that, when parents care for children through altruistic links, any marginal willingness to pay approach that adequately accounts for parents’ preferences will also account for the value of welfare gains to future generations.

In order to illustrate the relevance of the new dimensions introduced by our framework, we calibrate our model to U.S. data and calculate the value that the population alive in 1965 would attribute to the mortality reductions observed between 1965 and 1995. Our results imply that the welfare gain for an 18-year-old individual is between 65% and 180% higher than what would have been estimated if the welfare that young adults derive from their children was ignored. The aggregate social value is between 24% and 72% higher. These differences are due to benefits from child mortality reductions that are enjoyed by young adults and that could not have been estimated had we not incorporated altruism and fertility in the analysis. In addition, our calibrated model generates as by-products some interesting numbers never before calculated in the literature. For example, a function of parameters that can be roughly understood as the monetary value of the emotional loss from the death of a child varies between \$1.2 million and \$3.9 million.

Though applied to the analysis of changes in health, the paper touches on the more general question of how to deal with children and future generations in the evaluation of policy interventions. Since mortality risk at any age is a completely forward-looking measure, if no economic benefit from children is assumed, adults who survive childhood place zero value on mortality changes at earlier ages. As a consequence, in the standard framework, the contribution of changes in child mortality to the welfare of an adult is null, since generations are not linked either through altruistic behavior or through non-altruistic reasons such as ‘old-age security’ motives.

The link between generations considered in this paper is based on parental altruism. Altruism in human societies is certainly not restricted to the immediate family or even to direct descendants. Nevertheless, the case for incorporating parental altruism toward children in public policy evaluation seems far stronger than any other.² Not only does this approach incorporate the preferences of a significant share of the population that is not allowed to decide (children), but it also establishes an intergenerational link that ultimately accounts for the benefits accrued by all future generations. In addition, the task of assigning values to welfare gains experienced by children is transferred to those who are the children’s legal guardians and already decide for them in all relevant dimensions of life, namely, the parents. In reality, future generations cannot voice their concerns or reveal preferences via market behavior, and children, because of lack of maturity and dependence on parental care, are not allowed to legally decide. For these same reasons, contractual arrangements involving children—or possibly unborn individuals—cannot directly incorporate future benefits in current evaluations of policy interventions (see Becker and Murphy, 1988). As an alternative, our model allows for endogenous fertility and altruism in order to transfer the benefits of certain policies across different generations. In our theory, as in reality, parents are the ones who decide and are liable for their children.

Another way to link two successive generations would be to study non-altruistic parents whose fertility decisions are driven by egoistic reasons, such as old-age security motives (see Nugent, 1985; Ehrlich and Lui, 1991; Cigno, 1993, 2006, among others). Though

¹ The main contributions to the theoretical literature were made by Schelling (1968), Usher (1973), Arthur (1981), and Rosen (1988, 1994). Empirical analyses using this theoretical framework include Nordhaus (2003), Murphy and Topel (2003), Garrett (2001), Philipson and Soares (2005), Becker et al. (2005), Soares (2006), and Becker and Elias (2007).

² Cropper and Sussman (1988) and Rosen (1994) consider the problem of marginal willingness to pay for reductions in mortality when individuals leave bequests to a single descendant. Under certain circumstances, this may correspond to the incorporation of altruism. Some of the recent literature incorporates future populations by extending discount rates across different generations. In principle, intergenerational discount is different from time discount, so they should not be treated as the same concept. Bergstrom (1982, 2006) and Jones-Lee (1992) study altruism within the family and its effects on the value of life, but with no considerations for costly fertility decisions. Bergstrom (1982) shows that “pure altruism” produces no additional value if altruistic concerns are dealt with in the first order condition. Jones-Lee (1992) provides several extensions and generalizations. The case of fertility poses different problems because, in a marriage, as studied in Bergstrom (2006), the utility functions of husband and wife are well-defined and the formation of a family is voluntary. Not only are children self-produced by parents, but children do not have preferences accepted as relevant for valuation of mortality risk. In any case, none of these papers include the choice of number of children together with the value that parents attach to children, and therefore cannot be extended to adequately account for future generations. An alternative to the parental perspective is the adults-as-children valuation, where adults would assess the value of risk reductions that would have occurred when they were children. This approach, however, lacks theoretical foundations (see Scapecchi, 2006).

we do not pursue that approach in this paper, our emphasis on parental altruism should not be taken to suggest that non-altruistic motives are not important for understanding intergenerational welfare. We consider fertility choices in altruistic models following a long tradition in economic literature (Becker and Barro, 1988, or, more recently, Doepke and Zilibotti, 2005; Golosov et al., 2007), and see it as a stepping stone for exploring parental behavior under other plausible assumptions and more general settings.³ Also, old-age motives for fertility choice seem more important in historical settings, where savings or other forms of intergenerational transfers were not possible. In modern societies, children's costs are typically higher than the monetary returns they provide.

The literature on the economic valuation of child mortality reductions is not as developed as that regarding adults. While some empirical studies have examined possible differences of values between adults and children, most estimates assume exogenous fertility and egoistic agents (see Dickie and Gerking, 2006, for a review of the literature).⁴ The literature, however, has noted that factors like the presence and age of children, family size, and the timing of fertility decisions are important determinants of the willingness to pay for changes in mortality risk (Cameron et al., 2007; Dickie and Ulery, 2001).

Several of these applications have therefore stressed that the evaluation of policies affecting children requires knowledge of parental preferences (see, for example, Scapecchi, 2006; Carlin and Sandy, 1991; Mount et al., 2000; Dickie and Ulery, 2001; Dickie, 2005; Jenkins et al., 2001). Endogenous fertility decisions have also been suggested as a way to deal with evaluation of welfare involving future generations, since exogenous fertility implies a series of puzzling considerations (see Preston, 1993; Deaton, 2004; Viscusi, 2005). Under exogenous fertility, an unrestricted number of births is always advocated because it increases the aggregate number of years lived in society and, hence, social welfare. Also, changes in mortality immediately before or after birth cannot be adequately taken into account since there are almost no expenditures undertaken by parents up to that point. As a consequence, there is no theoretical basis to determine how future generations should be treated.⁵

Our framework shows that these puzzling considerations disappear once altruism and fertility are incorporated into traditional models. Generally, the same principle explored here can be extended to any other policy context in which government interventions have their costs or benefits spread out through different generations.

The structure of the remainder of the paper is outlined as follows. Section 2 starts by presenting a simple model illustrating the main implications of the incorporation of fertility and altruism into “value of life” calculations. It then describes the general version of the model and derives a formula for the social value of changes in survival functions. Section 3 illustrates the empirical relevance of our approach by calculating the value of the mortality reductions experienced by the U.S. population between 1965 and 1995. Section 4 concludes the paper with some general remarks.

2. A model of fertility and the value of life

This section illustrates the consequences of altruism and endogenous fertility for the valuation of life expectancy changes, and shows that they are intrinsically related to the value that parents attach to children and future generations. We construct a simplified example, as close as possible to the simplest model presented in Rosen (1988), in order to highlight the dimensions added to the problem and to compare our results to the previous literature. The generalizations required to make the model applicable to a quantitative study are discussed at the end of the section.

2.1. A simple model

Consider individuals who live for three periods: childhood, young adulthood, and mature adulthood. Individuals face a probability p_c of surviving birth. If they survive birth, they become young adults and face a probability p_a of survival into mature adulthood. Decisions are made at this stage, after individuals survive child mortality, and just before their adult mortality is realized.⁶ If individuals survive into mature adulthood, their consumption and fertility plans are carried out, and then their offspring face the child mortality risk.

³ As in Becker and Barro (1988) and Golosov et al. (2007), we consider intergenerational links as direct connections between potential people. Golosov et al. (2007) are concerned with efficiency when population is endogenous. Our model is complimentary to theirs. Parental altruism is also appealing because parents experience a significant grief from the loss of a child, and sometimes seem to be willing to pay more to reduce the probability of accidents for their children than for themselves. There is a large literature on income transfers within the family as a way to test for parental altruism with mixed results (see, for example, Altonji et al., 1997; Cox and Rank, 1992; Duflo, 2000).

⁴ Examples of valuations where non-fatality risks are the main concern include Agee and Crocker (1996), Dickie and Messman (2004), Joyce et al. (1989), Liu et al. (2000), and Viscusi et al. (1987). A broad review of the literature is available in Scapecchi (2006), who suggests that the value of children's health benefits is often higher than that of adults, though this is not always the case. In terms of valuations of reductions in child mortality risk, studies are often based on the purchase or use of safety-related devices (see Blomquist et al., 1996; Carlin and Sandy, 1991; Jenkins et al., 2001).

⁵ Valuations made by nineteenth-century courts after wrongful deaths of children were often given in terms of replacement of children's lost wages but became unfounded when child labor was abolished (Zelizer, 1994). Valuations in terms of parental investments in the upbringing of a child are also contradictory, especially for very young children, because, as Zelizer (1994, p.135) notes, they “would lead to the awkward conclusion that the average child had a negative worth and its death was a benefit for parents.”

⁶ Though this sequence of events may seem excessively artificial, it keeps things as simple as possible. The interpretation of the different stages becomes more natural in the generalized version of the model. For simplicity, we also ignore general equilibrium effects of changes in mortality and analyze an economy without production. With the exception of Arthur (1981), all papers in this literature have ignored general equilibrium implications. In Arthur (1981), however, all fertility decisions are exogenous. In our model, we address the economic value of changes in health that are outside the control of any particular individual (as related to scientific and technological developments in medical and biological sciences, for example). If households have control over some aspects of mortality, endogenous mortality changes will have no effect on welfare in addition to that reflected on expenditures.

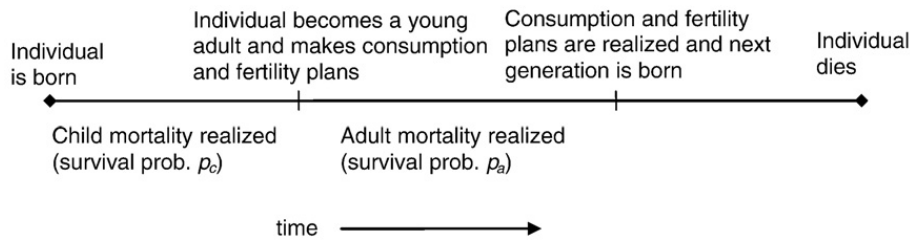


Fig. 1. Sequence of events in the simple model.

At time t , young adults receive an endowment w_t . Young adults decide on consumption and number of births before the event with probability p_a is realized. Actuarially fair insurance is available for every good consumed by parents, so that the budget constraint can be written as

$$w_t = p_a[c_t + n_t b + (p_c n_t) e], \tag{1}$$

where b is the goods cost of having a child and e is the goods cost of raising a surviving child. The costs of having and raising children are assumed to be fixed. To keep things simple and comparable to the previous literature, we assume no margin in which parents can invest or transfer additional resources to their children. In particular, we assume no intergenerational transfers as those typical of non-altruistic models of fertility choice or models of investments in children.⁷ These assumptions, however, are not crucial for our results and relaxing them would mainly suggest a different interpretation of the benefits of children in our model.⁸

Adults value consumption, the number of surviving children they have, and the utility that each child will enjoy as an adult. Adults are responsible for all decisions in the economy. Fig. 1 summarizes the sequence of events in this simple version of the model.

As in Rosen (1988), utility is assumed to be state-dependent in dimensions involving life and death events. We assume that surviving individuals derive utility from consumption and children, but that utility in case of death is equal to a constant M_a . Similarly, we assume that even though parents derive utility only from surviving children, there is a state-dependent utility loss associated with the death of a child. Therefore, the expected utility for a young adult satisfies the following value function

$$\tilde{V}_t = \max_{c_t, n_t} \left\{ p_a \left\{ \tilde{u}(c_t) + \mathbb{E} \left[\tilde{v}(n_t, N_t, \tilde{V}_{t+1}) \right] \right\} + (1-p_a) M_a \right\},$$

subject to the constraint (1), where c_t is consumption by adults at time t , n_t is the number of children born, N_t is the number of surviving children, and \tilde{V}_{t+1} is the future utility that the surviving children will enjoy as they become adults. The function $\tilde{u}(c_t)$ denotes the utility derived from consumption, and $\tilde{v}(n_t, N_t, V_{t+1})$ denotes the overall utility derived from children, including the enjoyment from the surviving children and their welfare, and the emotional loss associated with children who are born but do not survive childhood ($n_t - N_t$). Since there is an additional dimension of uncertainty related to the survival of children (probability p_c), parents try to maximize $\tilde{u}(c_t)$ plus the expected value of $\tilde{v}(n_t, N_t, V_{t+1})$, conditional on survival into adulthood.

We assume that the utility derived from surviving children follows the usual formulation from the fertility literature (as, for example, in Becker and Barro, 1988), so that the utility derived from N_t surviving children is given by $\alpha N_t^\phi V_{t+1}$, with $0 \leq \phi \leq 1$. As mentioned before, we also consider a state-dependent utility loss of αM_c for each child born who does not survive childhood. The utility loss associated with the death of $(n_t - N_t)$ children is given by $\alpha(n_t - N_t)M_c$. In our setting, M_c will have the connotation of the utility value associated with the loss of one child (regardless of family characteristics). This is a state-dependent formulation in the event of death of a child, analogous to the approach adopted by Rosen (1988) in relation to adult deaths. Finally, in order to simplify the analysis, we assume that there is no uncertainty about child quality or, in other words, that the series $\{w_i\}_{i=t}^\infty$ is known. Therefore, there is no uncertainty regarding the value of V_{t+1} itself.

In this context, we can normalize adult utility in the death state to zero by defining the instantaneous utility function $u(c_t) = \tilde{u}(c_t) - M_a$. By defining the value function $V_t = \tilde{V}_t - M_a$, we can specify a young adult's objective function as

$$V_t = \max_{c_t, n_t} \left\{ p_a \left\{ u(c_t) + \alpha \mathbb{E} \left[N_t^\phi V_{t+1} - (n_t - N_t) M_c \right] \right\} \right\}. \tag{2}$$

⁷ One way to introduce non-altruistic motives is to consider a family 'constitution' that specifies an amount of transfers between generations in the following way (see Cigno, 2006, for a complete treatment). Each adult in period t makes a transfer s_t to their parents and receives from each of their children, if both parents and children survive, a transfer of s_{t+1} . The budget constraint will then be given by:

$$w_t + p_a \left(\frac{s_{t+1}}{1+r} \right) (p_c n_t) = p_a [c_t + n_t b + (p_c n_t) e + s_t],$$

which is a generalization of our constraint in Eq. (1). In this case, there are incentives for intergenerational cooperation and children will be valued because the family complements the credit market. In what follows, we assume that transfers and investments in child quality are not allowed. As evident from the previous expressions, parental altruism and non-altruism motives entail their own sets of simplifications and assumptions.

⁸ Our assumptions also eliminate the traditional quantity-quality trade-off. Parents could still value child quality if we make the value function depend on human capital or bequests. Nevertheless, when the model is applied to data to evaluate different situations, a quantity-quality trade-off and non-altruistic motives will be reflected in different values of the parameters b and e , since these reflect the marginal value that parents attach to children.

Notice that there are two different dimensions of uncertainty in the problem. One is related to whether the individual will die before being able to realize the consumption and fertility plans, while the other is related to the number of children who survive (N_t) out of the total number of children born (n_t). Since the value function is defined as the expected utility of an adult individual, the only remaining dimension of uncertainty after the individual survives adult mortality is related to the number N_t of children who will survive out of n_t births.

This problem is as close as possible to the one discussed in Rosen (1988, Section 1), once fertility and altruism are incorporated in the analysis. Nevertheless, altruism and fertility introduce new, non-trivial dimensions of uncertainty in the discussion. Since one of our main motivations is to explore child mortality in a context of value of life calculations, we deal explicitly with the uncertainty related to p_c . Following Sah (1991), with constant p_c , the number of surviving children N_t follows a binomial distribution with density function

$$f(N_t, n_t) = \binom{n_t}{N_t} p_c^{N_t} (1-p_c)^{n_t-N_t}, \text{ for } N_t = 0, 1, 2, \dots, n_t.$$

Therefore,

$$\mathbb{E}[N_t^\phi V_{t+1} - (n_t - N_t)M_c] = V_{t+1} \sum_{N_t=0}^{n_t} N_t^\phi f(N_t, n_t) - (1-p_c)n_t M_c,$$

since $\mathbb{E}(n_t - N_t) = (1-p_c)n_t$ and V_{t+1} is known to parents.

As in Kalemli-Ozcan (2002), we approximate the function $\mathbb{E}[N_t^\phi]$ around the expected number of survivors $\mathbb{E}[N_t]$ using the delta method. This strategy allows us to deal with fertility as a continuous variable and still account for the effects associated with the risk regarding the number of surviving children via the explicit consideration of the second moment of the distribution. A second order approximation to $\mathbb{E}[N_t^\phi]$ leads to

$$\mathbb{E}[N_t^\phi] \approx (p_c n_t)^\phi - \frac{\phi(1-\phi)(1-p_c)(p_c n_t)^{\phi-1}}{2}, \tag{3}$$

since $\mathbb{E}[N_t - (p_c n_t)] = 0$ and the variance for the binomial distribution satisfies $\mathbb{E}[N_t - (p_c n_t)]^2 = n_t p_c (1-p_c)$. Using this result, the individual problem becomes

$$V_t = \max_{c_t, n_t} p_a \left\{ u(c_t) + \alpha \left[(p_c n_t)^\phi - \frac{\phi(1-\phi)(1-p_c)(p_c n_t)^{\phi-1}}{2} \right] V_{t+1} - \alpha(1-p_c)n_t M_c \right\},$$

subject to Eq. (1).

The first order conditions determining the optimal consumption and fertility decisions are given by the following expressions

$$p_a u_c(c_t) = p_a \lambda_t, \quad \text{and} \quad p_a \alpha \left\{ \frac{\partial \mathbb{E}[N_t^\phi]}{\partial n_t} V_{t+1} - M_c(1-p_c) \right\} = p_a (b + e p_c) \lambda_t,$$

where λ_t is the multiplier on the budget constraint and, from the calculations above,

$$\frac{\partial \mathbb{E}[N_t^\phi]}{\partial n_t} = \phi (p_c n_t)^{\phi-1} p_c + \frac{\phi(1-\phi)^2 (1-p_c)(p_c n_t)^{\phi-2} p_c}{2}.$$

Notice that since the variance of the number of survivors is a decreasing function of p_c (as long as $p_c \geq 1/2$), an increase in child survival tends to reduce the demand for children for risk averse parents. This effect, first considered by Sah (1991), takes place because the only form of “insurance” against child mortality risk is a large number of births (see Sah, 1991; Kalemli-Ozcan, 2002). On the opposite direction, since survivors are also less costly monetarily and emotionally when mortality declines, parents have an incentive for higher fertility in response to reductions in child mortality risk. The overall effect of child mortality on fertility is therefore ambiguous, although a reduction in mortality at low mortality levels would most likely reduce fertility rates (see Sah, 1991, p.591–592).

2.1.1. Changes in adult survival

The value of changes in life expectancy can be expressed as the marginal willingness to pay for increases in the probabilities p_a and p_c . For p_a , we define MWP_t^a as:

$$MWP_t^a = \frac{\partial V_t}{\partial p_a} \frac{1}{\lambda_t}.$$

This marginal willingness to pay can be interpreted as the monetary value of the marginal utility from increased survival probability, or, alternatively, as the marginal rate of substitution between survival probability and income. From the envelope theorem, this expression can be written as

$$MWP_t^a = \frac{u(c_t) + \alpha \mathbb{E}[N_t^\phi] V_{t+1} - \alpha M_c n_t (1-p_c)}{\lambda_t} + \frac{p_a \alpha \mathbb{E}[N_t^\phi]}{\lambda_t} \frac{\partial V_{t+1}}{\partial p_a} - c_t - b n_t - [p_c n_t] e.$$

Using the first order conditions, we obtain

$$MWP_t^a = \left(\frac{1}{\varepsilon_c} - 1\right) c_t + (b + ep_c) \left(\frac{1}{\varepsilon_n} - 1\right) n_t + p_a \alpha \mathbb{E}[N_t^\phi] \frac{\partial V_{t+1}}{\partial p_a} \frac{1}{\lambda_t},$$

where ε_c and ε_n denote the elasticities of the consumption and fertility sub-utility functions in relation to their respective arguments. In a stationary environment, we have that $\lambda_t = \lambda_{t+1} = \lambda$ and $\frac{\partial V_{t+1}}{\partial p_a} = \frac{\partial V_t}{\partial p_a} = \frac{\partial V}{\partial p_a}$. In this case, we can write

$$MWP^a = \frac{1}{1 - p_a \alpha \mathbb{E}[N^\phi]} \left[\left(\frac{1}{\varepsilon_c} - 1\right) c + \left(\frac{1}{\varepsilon_n} - 1\right) (b + ep_c) n \right]. \tag{4}$$

Eq. (4) can be immediately compared to the results from the original value of life literature. It highlights the insights gained by incorporating fertility and altruism in the analysis. Without fertility decisions and altruism, Eq. (4) would be reduced to $(\varepsilon_c^{-1} - 1)c$, which is exactly the result presented in Rosen (1988, p.287). This result summarizes the fact that increases in life expectancy will be more valuable for higher levels of consumption and for a lower elasticity of the sub-utility function $u(c)$. However, because higher survival probabilities increase the cost of the actuarially fair insurance, reductions in mortality lower consumption in case of survival. This is precisely the sense in which Rosen (1994) identifies a trade-off between the quantity and the quality of life.

The second term inside brackets represents an analogous effect in relation to fertility, coming from the fact that increased survival probability increases the likelihood that individuals will live long enough to realize their fertility plans and derive utility from their children. The term $(b + ep_c)$ converts the value of fertility into monetary (consumption) units, while the rest of the expression is completely analogous to the case of consumption. The new expression states that increases in life expectancy will be more valuable in high fertility societies. Alternatively, Eq. (4) implies that willingness to pay for changes in adult mortality risk depend on parenthood status and the number of children in the household (see Cameron et al., 2007; Jenkins et al., 2001, for empirical estimates of the role of fertility in adult willingness to pay). As in the case of the consumption factor, increases in life expectancy are more valuable for low values of the elasticity of the fertility sub-utility in relation to its argument. This effect derives from the fact that increases in survival probabilities increase the cost of survivors, therefore reducing fertility in case of survival.

Expression (4) highlights the key roles played by the elasticities of the sub-utility functions in relation to c and n . If these elasticities were the same, it would make no difference to separate expenditures on children from those on consumption, since $\left(\frac{1}{\varepsilon_c} - 1\right)c + \left(\frac{1}{\varepsilon_n} - 1\right)(b + ep_c)n$ would be equal to $\left(\frac{1}{\varepsilon_c} - 1\right)[c + (b + ep_c)n]$. In other words, from the perspective of changes in adult mortality, it would be unnecessary to take into account explicitly expenditures on children as long as $\varepsilon_c = \varepsilon_n$. It is the difference between these elasticities that distinguishes the gain in welfare from the increased probability of future consumption, from the gain in welfare from the increased probability of having children.

Finally, the term outside the brackets adjusts the value of life for the fact that not only will these changes affect the current generation, but they will also benefit all future generations. MWP_a accounts for the present discounted value of the welfare gains: $[(\varepsilon_c^{-1} - 1)c + (b + ep_c)(\varepsilon_n^{-1} - 1)n]$ per generation, discounted at a rate $p_a \alpha \mathbb{E}[N^\phi]$.⁹ In order to write the present value of welfare gains for all future generations in an expression as simple as the one above, we assume a stationary environment.

Our methodology therefore evaluates the welfare gain from certain changes in mortality rates if current conditions were maintained indefinitely into the future. In other words, our framework tells us what the value of certain life expectancy improvements would be in a stationary world where the conditions observed in the data persisted forever. To the extent that economies display long-run growth, it will underestimate the true welfare improvements brought about by reductions in mortality rates. Even with these limitations, our setup extends the analysis in a direction that was entirely ignored by the previous literature.

The results discussed in this section neatly illustrate the features incorporated in the valuation of adult mortality changes once fertility and altruism are taken into account. Apart from the usual effects, a permanent decline in adult mortality benefits current and future generations due to altruistic links, and to the higher probability that adults will live long enough to have children. In reality, the establishment of an altruistic link between parents and children ends up incorporating the welfare gains of all future generations. In our framework, the weight attributed to these future generations is directly related to the value that current adults—the actual decision-makers—attribute to their own children.

⁹ For this to be the case, we must have $p_a \alpha \mathbb{E}[N^\phi] < 1$, which is precisely the condition required for the recursive problem described before to be well defined. Note also that the homogeneity of the value function, characteristic of the deterministic version of Becker and Barro (1988), fails to hold in the presence of child mortality under our state-dependent formulation. Notice also that our assumption of an exogenous value V_{t+1} can be relaxed by allowing for bequests without major changes in the results. If B_{t+1} is the optimal level of bequests parents leave to each surviving child, under endogenous bequests, the value function becomes $V(B_{t+1})$. Since decisions about bequests are dealt with in the first order condition, the expressions for MWP will not change although, as we have previously noted, the interpretation given to the costs of children will be different.

In addition to these gains, there are important aspects of changes in child mortality and their interactions with fertility that are distinct in nature from the effects of changes in adult mortality. We now explore these implications within the framework of our model.

2.1.2. Changes in child survival

We define MWP^c as the marginal willingness to pay of adults for increases in p_c . As we stated before, parents, not children, express their concern for survival and their willingness to pay for mortality changes. From the envelope theorem and using the first order conditions, we can write MWP^c as

$$MWP_t^c = \frac{p_a}{p_c} \left\{ bn_t + \frac{(1-\phi)}{2 + (1-\phi)^2 \mu^2} (b + p_c e) \right\} + p_a \alpha \left\{ \mathbb{E}[N_t^\phi] \frac{\partial V_{t+1}}{\partial p_c} + M_c n_t \right\} \frac{1}{\lambda_t},$$

with $\mu = \sqrt{\text{Var}(N)}/\mathbb{E}[N]$ as the coefficient of variation.

Again, in a stationary environment we can write

$$MWP^c = \frac{1}{1 - p_a \alpha \mathbb{E}[N^\phi]} \frac{p_a}{p_c} \left[bn + \frac{(1-\phi)}{2 + (1-\phi)^2 \mu^2} (b + p_c e) + p_c \frac{\alpha M_c}{u_c(c)} n \right]. \tag{5}$$

The term transforming the flow of welfare into its present discounted value multiplies the expression as in the marginal willingness to pay for changes in adult survival rates. Changes in child mortality are assumed to be permanent, and, therefore, they will be enjoyed not only by this generation but also by all future generations.

In the case of child mortality, three elements compose the valuation of mortality changes for any given generation (the expression multiplying the discount factor). The first term inside square brackets (bn) represents savings for households due to lower costs in the acquisition of survivors, while the second term derives exclusively from uncertainty considerations. The last term represents the increased welfare from the reduced probability of death of a child (αM_c per child monetized by $u_c(c)$).

Three extreme examples help clarify the economic forces at work here. First, with risk neutrality ($\phi=1$), the second term inside brackets disappears. In this case, the goods cost per child born and the utility cost per child death determine the marginal willingness to pay for reductions in child mortality. Risk neutrality in this context can also be interpreted as perfect replacement in fertility, since parents have no incentive for fertility insurance or hoarding motives. One can think of it as if children could be replaced immediately after they died, and there was no need to decide on the number of children before child mortality was realized.¹⁰

Second, consider the other extreme when $0 < \phi < 1$ and $b = M_c = 0$. Here, the goods cost of children is defined only for survivors, and there is no welfare loss due to the death of a child. Hence, the value of changes in child mortality depends only on the risk premium due to uncertainty. In this case, even without additional economic costs, reductions in child mortality may improve welfare since increased chances of child survival reduce parents' uncertainty.

Finally, if we have $\phi = 1$, still with $M_c = b = 0$, there are no gains from reductions in child mortality. In this situation, the number of births adjusts on a one-to-one basis to changes in child mortality, and all costs of children are defined in terms of survivors. As parents are risk neutral and there is no cost wedge between children born and surviving children, parents have a target number of expected survivors that is maintained irrespective of the child mortality rate.

In summary, two elements compose the valuation of reductions in child mortality: costs of non-surviving children (both monetary and utility losses) and reductions in the uncertainty associated with surviving children. These gains are different in nature from the traditional valuations in Arthur (1981) and Rosen (1988), and deserve particular attention in an attempt to attach social values to health improvements. In our setup, the willingness to pay for changes in child mortality is not expressed by the individuals who are subject to the mortality risk, since these individuals (children) are not legal decision-makers and, in some circumstances, have not yet been born. This is precisely why a framework that adequately accounts for the value of changes in child mortality must incorporate fertility as a choice variable, and must express welfare through the perspective of parents and their altruism toward children.

2.2. A general model

This sub-section generalizes the model presented before. Our goal is to obtain expressions that allow the valuation of specific mortality changes for an adult individual at any given age. With that in hand, the model can be used to determine the social value of any change in survival probabilities.

As will be clear below, a critical variable for the calculation of the social value of mortality changes is the value of these changes for an individual entering adulthood. An analogous expression to Eq. (2) for a representative individual entering adulthood at age a is:

$$V_T^a = \int_a^\infty u(c(t))S(t, a)dt + \alpha S(\tau, a) \left\{ \mathbb{E}[N_T^\phi] V_{T+1}^a - (1-p_c)n_T M_c \right\},$$

¹⁰ Estimates of the extent of behavioral replacement effects are often below one with point estimates near or below one-half. That is, additional births do not completely replace child deaths in the family. See Wolpin (1997, Section 2.5) for a survey of empirical results. Since parents might have already anticipated a child death, and insured against it, it is natural to assume that incomplete replacements are simply a consequence of anticipatory fertility choices that call for no additional births.

where $S(t, a)$ is the discounted survivor function, or the function describing the probability that the agent survives from ages a to t , discounted at the rate of time preference.¹¹ At age a , individuals decide their profile of consumption, and at age $\tau > a$, parents realize their fertility plans. All children are assumed to be born at the same time τ . As before, V_{T+1}^a reflects the children's utility once they reach adulthood at age a . T indexes generations and t indexes ages. N_T represents the number of children surviving to age a out of a total number of births n_T .

For the sake of simplicity, we assume that children are born, face all the risks related to child and pre-adult mortality at one single moment, and immediately become adults. Surviving children are binomially distributed as in Eq. (3). In terms of the undiscounted survival function $S^*(t, i)$, we only consider the final survival probability between ages 0 and a , $p_c = S^*(a, 0)$, where, as before, p_c represents the total survival probability for children. We revisit this issue when calibrating the model (see Section 3). The rest of the model remains unchanged. For a given endowment W_T received by a member of generation T at age a , households have the following budget constraint

$$W_T = \int_a^\infty c(t)S(t, a)dt + S(\tau, a)(bn_T + e[p_cn_T]),$$

where the interest rate is assumed to be equal to the subjective discount rate.

This formulation keeps the basic features discussed before, but adds a couple of new dimensions in terms of the impact of adult mortality reductions across different age groups. For example, adult mortality reductions taking place before fertility decisions are realized (τ) have now qualitatively different impacts from mortality reductions taking place after that. Or, alternatively, adult mortality changes related to ages below that of parents, which therefore do not affect their personal survival, may still affect their welfare through intergenerational links.

First order conditions in this case are almost identical to those discussed in the simpler version of the model. To evaluate changes in mortality in this context, we think of shifts in the survival function as caused by changes in some underlying parameter θ . In this sense, we define $S_\theta(t, i) = \partial S(t, i; \theta) / \partial \theta$ as the change in the conditional discounted probability of survival from age i to age t brought about by a change in θ .

2.2.1. Changes in survival for an individual entering adulthood

We start with the problem of an individual entering adulthood, at age a . We label the willingness to pay of generation T at age a for changes in adult survival as MWP_a^A . Following the steps carried out above, one can show that in a stationary environment we have

$$MWP_a^A = \frac{1}{1 - \alpha S(\tau, a) \mathbb{E}[N^\phi]} \left[\int_a^\infty \left(\frac{1}{\varepsilon_c} - 1 \right) c(t) S_\theta(t, a) dt + S_\theta(\tau, a) (b + p_c e) \left(\frac{1}{\varepsilon_n} - 1 \right) n \right]. \quad (6)$$

When mortality reductions affect survival probabilities before age τ , individuals also benefit directly from the increased probability of living long enough to have children. Since individuals are altruistic, they are willing to pay to ensure new generations through the realization of fertility decisions. This same force is not present when changes in mortality only affect survival rates after age τ .

The case of changes in child survival is completely analogous to the simpler version of the model. We define MWP_a^C as the marginal willingness to pay of an adult at age a for increases in p_c . In a stationary environment, we obtain an expression almost identical to Eq. (5):

$$MWP_a^C = \frac{1}{1 - \alpha S(\tau, a) \mathbb{E}[N^\phi]} \frac{S(\tau, a)}{p_c} \left[bn + \frac{(1 - \phi)}{2 + (1 - \phi)^2 \mu^2} (b + p_c e) + \frac{\alpha p_c M_c}{u_c(c)} n \right] \frac{\partial p_c}{\partial \theta}. \quad (7)$$

The interpretation of the different terms in the equation above was already discussed above.

2.2.2. Changes in survival for a young adult

Define young adults as individuals who have not yet acted on their fertility choices. For these individuals at age i , where $a \leq i < \tau$, the questions involved are similar to the ones already discussed but for one small detail. Since individuals will not necessarily be at age a , we will not be able to isolate the marginal change in the value function on the left hand-side in order to obtain a simple expression for the marginal willingness to pay, as we did before.¹² This is exactly why MWP_a^C and MWP_a^A play such critical roles.

¹¹ If $S^*(t, a)$ is the survival function, the discounted survival function is given by $S(t, a) = e^{-\rho(t-a)} S^*(t, a)$, where ρ is the subjective discount rate. Strictly, consumption should also be indexed by generation, as in $c_T(t)$, indicating the consumption of generation T at age t . We use the discounted survival function and write $c(t)$ to save on notation.

¹² The marginal willingness to pay refers to an individual at age i , while the derivative of the value function refers to an individual at age a . Therefore, we cannot transfer the derivative of the value function to the left hand-side in order to obtain the present discounted value of changes in mortality for all future generations. However, since the solution to the individual problem of a young adult is time consistent, the multiplier on the budget constraint will always have the same value in a stationary environment, irrespective of generation or age (this can be seen from the first order conditions and the budget constraint). In a stationary environment, this means that the marginal utility of wealth (or the multiplier) will be constant within and across generations.

In a stationary environment, the marginal willingness to pay of an individual at age i for changes in adult survival rates is

$$MWP_i^A = \int_i^\infty \left(\frac{1}{\varepsilon_c} - 1\right) c(t) S_\theta(t, i) dt + S_\theta(\tau, i) (b + p_c e) \left(\frac{1}{\varepsilon_n} - 1\right) n_T + \alpha S(\tau, i) \mathbb{E}[N_T^\phi] MWP_a^A.$$

As this expression makes clear, MWP_a^A , or the marginal willingness to pay of an individual at age a , is a key concept in the problem. Our willingness to pay measure allows the valuation of future generations to be incorporated via the altruism of parents. In this case, the willingness to pay of parents at age i is composed, as before, by the consumption and fertility factors, plus the discounted value of these gains for their children. In turn, children will become adults when their parents are τ years old. The meaning of this expression is exactly the same as Eq. (6), the only difference being that parents' marginal willingness to pay will not be exactly the same as their children's. Therefore, it is not possible to express MWP_i^A directly as the simple discounted value of the sum of the consumption and fertility factors.

For changes in θ that affect child mortality, time consistency of the individual's problem allows us to write

$$MWP_i^C = \frac{S(\tau, i)}{p_c} \left\{ \left(b + \frac{\alpha p_c M_c}{u_c(c(t))} \right) n_T + \frac{(1-\phi)}{2 + (1-\phi)^2 \mu^2} (b + p_c e) \right\} \frac{\partial p_c}{\partial \theta} + \alpha S(\tau, i) \mathbb{E}[N_T^\phi] MWP_i^C.$$

This expression has the same interpretation as Eq. (7). The only difference is that altruism for future generations is expressed via the discounted value of the marginal willingness to pay of individuals entering adulthood (MWP_a^C), and not of individuals at the same age as their parents.

2.2.3. Changes in survival for an old adult

For those individuals at age $i > \tau$, fertility decisions have already been realized. As mentioned before, we assume that these individuals cease to be altruistic toward their offspring. With the hypothesis of egoistic old adults, the problem of an individual at age $i > \tau$ turns into the classical problem of the value of life literature. If mortality reductions take place in age groups that do not directly benefit the individual, the marginal willingness to pay is zero. We define the marginal willingness to pay of an old adult for changes in θ as MWP_i^O , and write:

$$MWP_i^O = \int_i^\infty \left(\frac{1}{\varepsilon_c} - 1\right) c(t) S_\theta(t, a) dt. \tag{8}$$

This corresponds to what we called before the consumption factor. It represents the goods trade-off between quantity and quality of life identified by Rosen (1994).¹³

2.2.4. Social value of mortality reductions

With the previous results in hand, we are able to determine the social value of any given change in the survival function. The social value will be a function of the age distribution of the population, and of the valuation attached to the changes by each specific age group. The valuation of each age group will ultimately depend on the moment and extent of the observed reductions in mortality.

Suppose that a population of P individuals is distributed across ages according to the distribution function $F(\cdot)$. The social value of a given change in mortality rates, brought about by a shift in the parameter θ , is given by

$$\text{Social MWP} = P \left[\int_a^\tau (MWP_i^C + MWP_i^A) dF(i) + \int_\tau^\infty MWP_i^O dF(i) \right].$$

Or, defining

$$MWP_i = \begin{cases} MWP_i^C + MWP_i^A, & \text{if } i \leq \tau \text{ and} \\ MWP_i^O, & \text{if } i > \tau \end{cases}$$

we can write it in a simpler form

$$\text{Social MWP} = P \int_a^\infty MWP_i dF(i). \tag{9}$$

The social value is simply the weighted sum of the value of mortality changes across the different age groups, where the weights are given by the number of individuals in the population that belong to each specific group ($PdF(i)$). It is immediate to see that the age distribution may have a significant impact on the social value of a specific change in survival probabilities. A society with a high proportion of old individuals will value reductions in child or young adult mortalities less, and value extensions in old age life expectancy more.

¹³ Though certainly extreme and concentrated at one point in time, the distinction between young and old adults captures the idea that the value that parents place on a child's life is reduced as the child ages. This pattern has been noticed in the empirical literature by Jenkins et al. (2001). The value that parents place on adult mortality risks also varies with parenthood status and the number of children (see Cameron et al., 2007).

Also notice that our aggregate willingness to pay only incorporates individuals at age a or older. From a policy perspective, these are the actual decision-makers and the individuals who incur the costs of child raising. Legally, they are the ones who are allowed to decide and who can interfere directly in the political process. Children and future generations only enter the social willingness to pay via the altruism of current adults. This structure captures in a stylized way the actual process of political decision observed in modern societies and avoids the double-counting of value present in settings where altruism is important (Bergstrom, 1982; Jones-Lee, 1992).

In reality, Eq. (9) is analogous to the characterization of the optimal provision of a public good common in the public finance literature. Eq. (9) gives the maximum amount of taxes that should be raised on the current population in order to promote a given reduction in mortality rates (a change in θ). The key difference here is that our measure of social willingness to pay incorporates the gains of children and future generations. These are considered to the extent that current adults—due to the links established by intergenerational altruism—are willing to pay for the welfare gains of their descendants. Indeed, Eq. (9) characterizes the optimal allocation from the perspective of the costs and benefits faced by the population currently alive.¹⁴

3. Parametrization and empirical implications

3.1. Calibration

In order to illustrate the importance of the dimensions added by our framework, we use the model developed in the previous section to give monetary values to the mortality reductions experienced by the United States population between 1965 and 1995. Though part of the improvements in health during this period were endogenous to changes in habits and income, our goal is simply to use this example to show how important fertility and altruism can be in assessing the economic value of reductions in mortality. This is a case that has been studied in the literature and, therefore, constitutes a good initial example. In addition, the United States is one of few countries for which alternative estimates of the “value of a statistical life” for both adults and children are currently available.¹⁵

We look at the problem from the perspective of individuals alive in 1965, and ask how much they would be willing to pay for the gains in survival probabilities that actually took place in the following 30 years. As discussed previously, the numbers generated by the methodology assume that individuals consider the economic conditions prevalent in 1965 as the ones that would persist indefinitely into the future. As the marginal willingness to pay increases with income, the perspective of economic growth would increase the value of mortality reductions and even more so for future generations. So, as with many other dimensions of our empirical strategy, this assumption tends to reduce both the value of life expectancy gains and the relative importance of altruism and fertility. As a general rule, whenever faced with choices regarding the implementation of our methodology, we choose to proceed with the option that minimizes the importance of the very factors we want to highlight. In this manner, we hope to guarantee that any remaining bias will work against the relevance of our main points. To proceed, we need values for a series of parameters and variables determining initial conditions (c , n_T , b , e , ϕ , α , and the interest rate), the survival function ($S^*(t, a)$), and the change in the survival function between 1965 and 1995 ($S_{\theta}^*(t, a)$).

The U.S. survival functions for 1965 and 1995 are calculated from age-specific number of deaths and population, obtained from the World Health Organization Mortality Database. The WHO database contains information broken down by 5-year age intervals, so we assume constant mortality rates within each interval. We calculate the 1965 numbers as averages for the period between 1960 and 1969, and the 1995 numbers as averages between 1990 and 1999. Fig. 2 illustrates the shift in the (undiscounted) survival function for the U.S. population between 1965 and 1995. This shift corresponds to a change in overall life expectancy at birth from 70 to 76 years. Note that the U.S. experienced relatively modest gains in child mortality during this period, mainly because of the already extremely low starting point (mortality before age 5 went from 2.84% to 0.99%). Also, the age distribution of the American population in 1965 had a relatively high share of old adults when compared to young adults (36% of the population above age 40). If anything, the example chosen tends to minimize the importance of altruism, fertility, and future generations in the calculation of the social value of mortality reductions.

Of the parameters and variables needed, some are common in the “value of life” literature. First, interest rates are assumed to be 3% per year, and this number is used to calculate the discounted survival functions ($S(t, a)$). The assumption of subjective rate of time preference equal to the interest rate implies constant consumption throughout life. Since we are not interested in life-cycle considerations, and this facilitates the empirical implementation of the model, it is the best way to proceed. This allows the consumption term in the willingness to pay expressions to be factored as $(\delta_c^{-1} - 1) c_i^{\infty} S_{\theta}(t, i) dt$. Following Becker et al. (2005), we assume that the instantaneous utility function for consumption takes the form $u(c) = \varphi + c^{\sigma}/\sigma$, which already incorporates the

¹⁴ This is very different from the approach taken by Murphy and Topel (2003) or Soares (2006). Murphy and Topel (2003) and Soares (2006) incorporate future generations by simply extrapolating current numbers and discounting future populations at the interest rate. This amounts to assuming a specific type of intergenerational welfare function that, as every other welfare function, is arbitrary. For example, given the preferences they assume, their numbers would not be supported by voters as policy benchmarks for the benefits of improvements in health. Still, our specification only includes future generations to the extent that current generations are willing to pay something for their welfare. For example, if it were possible to finance expenditures on health by issuing debt to be paid by future generations, the optimal allocation would be different from that dictated by our willingness to pay approach.

¹⁵ The changes observed during the period are quite substantial. But, given the complexity of the problem at hand, closed form solutions cannot be obtained for the valuation of discrete changes in survival probabilities. Therefore, we stick to the formulas derived before for the marginal willingness to pay, and see our results as approximations of the true welfare value of mortality reductions.

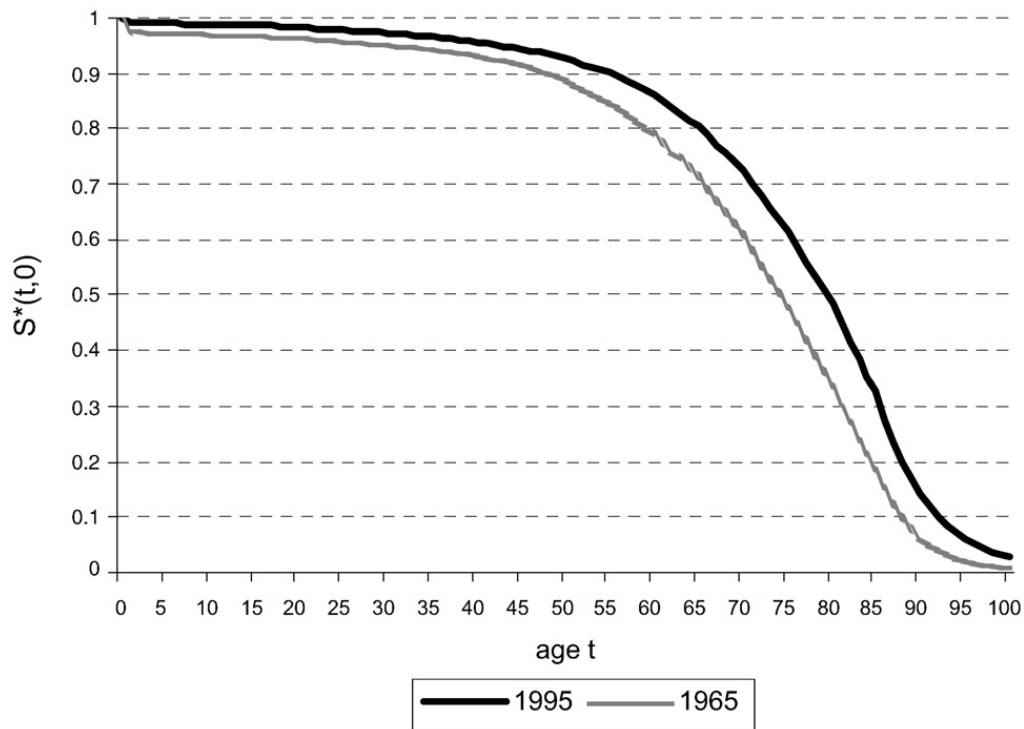


Fig. 2. Survival probabilities for the U.S. population (from age 0 to t), 1965 and 1995.

normalization of utility of death to zero (as in the formulation proposed by Rosen, 1988). We also stick to the calibration proposed by these authors, where $\varphi = -16.16$ and $\sigma = 0.2$.¹⁶

For fertility (n_T), the survival probability of children (p_c), and the discounted survival probability of adults up to the age when fertility is realized ($S(\tau, a)$), we use the 1965 values of, respectively, the total fertility rate (2.9, or $n_T = 1.45$, from the World Development Indicators, WDI), the survival probability between ages zero and 18 (0.96), and the discounted survival probability between ages 18 and 40 (0.50). We take p_c to be the survival probability from zero to eighteen because 18 years of age seem to be the natural benchmark for legal independence of children. Besides, this timing matches well the data on the costs of raising children, which will be used to calibrate some of the remaining parameters. In relation to the age when fertility is realized, we choose $\tau = 40$. Since we are consolidating the entire child-rearing process in a single moment, we choose the middle point of the parenthood process as such moment. If one thinks that parents usually start having children when they are between 25 and 30, and take care of children in one way or another until they are roughly 50 or 55, the middle point would be around age 40.

Estimates of the costs of raising a child (b and e) can be obtained from surveys of expenditures on children, such as the ones undertaken by the U.S. Department of Agriculture (USDA). Lino (2001) presents the 1960 estimates of the USDA for the average costs of raising a child up to age 18, in 2000 US\$. We use the CPI to deflate this value to 1996 US\$. The USDA estimates imply that roughly 5% of the expenditures on children takes place before age 2, while 95% takes place between ages 2 and 18. We assume that costs up to age 2 are related to the costs of having a child (b), and costs between ages 2 and 18 are related to costs of raising a surviving child (e). This yields $b = \$6956$ and $e = \$126,812$. If anything, these estimates seem to underestimate the value of b in comparison to e , and this will tend to minimize the importance of child mortality in welfare evaluations.

Consumption is set to the 1965 value of total consumption per adult (in 1996 US\$), net of expenditures on children. This value amounts to \$14,202, and corresponds to total private plus government consumption divided by the population above 18, minus yearly expenditures on children (the flow value of the unit cost of children times n_T).

The remaining parameters needed— α , ϕ , and M_c —are uncommon in applied work and, as far as we understand, have never been directly estimated in the literature. Therefore, some exploratory calibration effort is required, involving a subtler type of information that is not so widely available. Calibration of these parameters requires knowledge of how much individuals value their own survival probabilities vis-à-vis their children's. There is a small empirical literature, based on interpersonal willingness to pay for changes in mortality, that addresses this type of question.

Cropper et al. (1992), for example, use a survey to estimate the rate at which people implicitly discount future lives saved. But their questions do not imply any kind of personal relationship among the individuals being considered and, therefore, are not

¹⁶ Murphy and Topel (2003) incorporate leisure in the analysis and obtain a measure of full-income. Therefore, their numbers are larger than the ones obtained here. Murphy and Topel (2003) also deal explicitly with the issue of life-cycle changes in earnings and consumption. To keep things simple and to focus on the points of interest, we abstract from these considerations.

Table 1
Benchmark variables for base year (1965) and for calibration (1995)

Variable	Definition	Source	1965	1995
p_c	Survival probability to age 18	WHO	0.96	0.98
$S(\tau, a)$	Discounted survival prob. from age 18 to 40	WHO	0.50	0.51
c	Consumption per adult, net of expenditures on children	FED	14,202	27,738
n_T	Fertility rate per parent	WDI	1.47	1.02
b	Child cost before age 2	USDA	6,956	7,847
e	Child cost from age 2 to 18	USDA	126,812	143,052

Note: Values of c , b , and e in 1996 \$. Sources and construction of variables described in detail in the text. Values in 1995 used to calibrate α , ϕ , and M_c ; values in 1965 used as benchmark to calculate the monetary value of the reductions in mortality observed between 1965 and 1995.

adequate for our analysis. Carlin and Sandy (1991) use car safety-seat information to estimate the “value of a statistical life” of children to their mothers, but do not present comparable numbers for adults. Dickie and Ulery (2001) use a contingent valuation survey to estimate parents' willingness to pay for various health conditions affecting themselves and their children. These authors find that parents are willing to pay more for their children's health than for their own (willingness to pay to avoid acute illnesses), but do not address directly the issue of mortality.¹⁷

Three papers use information on adult decisions to estimate the “value of a statistical life” for both adults and children, and therefore contain the ideal type of information needed to calibrate our parameters of interest. Blomquist et al. (1996) use a set of different types of data (seat-belt use, child safety-seat use, and motorcycle helmet use) and obtain a wide range of estimates for adults and children. Their preferred estimate for the “value of a statistical life” for adults is \$2 million, while the average estimate for children is around \$4 million.

Mount et al. (2000) calculate “value of a statistical life” based on cars owned by different types of families (single or with children) and the safety features of the cars. From the specifications presented by the authors, our preferred value is the one that controls for differences in family income, leading to “value of a statistical life” of \$2.62 million for children and \$3.88 million for adults.

Finally, Jenkins et al. (2001) make use of data from purchases of bicycle helmets, which are available specifically for people at different age ranges and protect one particular individual within the family. From the estimates presented by the authors, our preferred value is the one that assumes that helmets are worn less than 100% of the time and that only half of the benefit comes from mortality reductions (the rest being attributed to morbidity). This specification leads to “value of a statistical life” of \$1.3 million for children and \$2 million for adults.

All these papers estimate the “value of a statistical life” for children and adults separately. These reflect values implicit in decisions taken by adults, matching well our theoretical formalization.

From the theoretical perspective, the “value of a statistical life” can be conceptualized as the marginal willingness to pay for a one-shot unit change in the probability of death at some specific age (as in Murphy and Topel, 2003). Therefore, from Eqs. (6) and (7), a given set of estimates of the “value of a statistical life” for adults and children defines—together with the variables discussed before—two non-linear equations on three unknown parameters (α , ϕ , and M_c). The first order conditions for c_t and n_T can be used to define a third equation, which, together with the previous two, defines a system of three equations on three unknowns.

Since the studies discussed use data from the 1990's, we set the other variables to their 1995 values in order to solve the system of non-linear equations for α , ϕ , and M_c . The Appendix A presents the system of three non-linear equations and discusses in detail how this calibration was conducted. For each set of estimates of the “value of a statistical life” for adults and children, we obtain a set of estimates of α , ϕ , and M_c . Table 1 presents the values used in the calibration of the parameters and in the calculation of the marginal willingness to pay for the mortality reductions observed between 1965 and 1995. The calibrated values of α , ϕ , and M_c are shown in Table 2.

In order to give a concrete meaning to the parameters presented in Table 2, we can multiply α by M_c and divide it by the marginal utility of consumption in order to get, roughly speaking, the monetary value of the emotional loss associated with the death of a child. The set of parameters calibrated imply that this value is between \$1.2 million and \$3.9 million. This is at least one order of magnitude greater than the utility that parents derive from actually having a child, which, from the first order conditions, must be closely related to its marginal cost (the utility from having a child can be expressed, in monetary terms, as $\alpha V_{t+1}/u_c(c)$, which revolves around \$100,000 in the specifications). According to any of the calibrated versions of the model, the utility loss associated with the death of an existing child is orders of magnitude greater than the utility gain from having an additional child.

3.2. Results

Table 3 presents the results. It contains the monetary value of child and adult mortality reductions for an individual entering adulthood (age 18) and for the entire 1965 American population. The table shows that any set of parameters yields a significant

¹⁷ Additional valuations for non-fatal outcomes are discussed in footnote 5. See also Scapecchi (2006) for a survey on the empirical findings on the economic valuation associated with reduced health risks to children.

Table 2Estimates of α , ϕ and M_c based on the “value of a statistical life” (VSL) for adults and children

Source of VSL	VSL (Mill., 1996 \$):		Calibrated values		
	Children	Adults	α	ϕ	M_c
Blomquist et al. (1996)	4.00	2.00	0.0542	0.5803	20,091
Mount et al. (2000)	2.62	3.88	0.0405	0.4762	17,464
Jenkins et al. (2001)	1.30	2.00	0.0303	0.3949	11,306

Note: Calibrated values from authors' calculations using the “value of a statistical life” estimated by Blomquist et al. (1996), Mount et al. (2000), and Jenkins et al. (2001). Calculations based on the three equations described in the Appendix A and the values of variables in 1995 presented in Table 1.

effect of changes in child mortality on the welfare of a young adult. The results imply that an 18-year-old individual in 1965 would be willing to pay between \$11,697 and \$36,449 for the changes in child mortality observed in the following 30 years, and between \$17,861 and \$20,536 for the changes in adult mortality observed in the same period. Reflecting the differences in the “value of a statistical life” used in the calibration, the different sets of parameters lead to differences in the relative importance of reductions in child and adult mortalities. In any case, all sets of estimates lead to welfare gains from reductions in child mortality that are of the same order of magnitude as those from reductions in adult mortality. With the Jenkins et al. (2001) numbers, the total welfare gain is 65% higher than what would be estimated by looking at reductions in adult mortality alone, while this difference jumps to 115% in the calibration based on Mount et al. (2000), and 180% in the calibration based on Blomquist et al. (1996). These are benefits from child mortality reductions that are enjoyed by young adults and that could not be estimated in the traditional set up.¹⁸

These different numbers reflect, to some extent, the different calibrated values of ε_n in relation to ε_c . The parameter ε_c is calibrated to 0.38 in all exercises, while ε_n varies from 0.38 using the Mount et al. (2000) numbers, to 0.57 using Blomquist et al. (1996), and to a negative -0.44 using the Jenkins et al. (2001) estimates. In the case of the parameters calibrated from the Mount et al. (2000) numbers, ε_n is virtually identical to ε_c . But more surprising is the case of the calibration with the Jenkins et al. (2001) estimates, where ε_n reaches a negative value. This would correspond to the implausible situation where parents derive negative utility from having children, but still are willing to pay positive amounts to reduce child mortality because of the emotional loss associated with the death of a child. In this particular case, the model would not be able to explain why parents do have children. This strange result comes from the very low value of a statistical life for children—both in absolute and relative terms—estimated in the Jenkins et al. (2001) study (\$1.3 million, as compared to \$2 million for adults). Still, in all cases, our estimated value for changes in adult mortality (MWP_a^A) is not very much different from what would be obtained using the traditional framework. This means that the distinction between gains associated with the higher probability of living longer to consume more and the higher probability of living long enough to have children does not seem particularly important. The real quantitative difference arises from the capacity of our model to express the willingness to pay of adult individuals for reductions in child mortality (MWP_a^C), and from the fact that this willingness to pay turns out to be quite large, independently of the set of parameters used.

When we move to the analysis of the social value of mortality changes, the relative importance of child mortality is somewhat reduced, since a large fraction of the 1965 population was composed of old adults who had already surpassed the assumed child-rearing age ($\tau=40$). Still, Table 3 shows that the social value estimated by our methodology is between 24% and 72% higher than what would be obtained if we looked at adult mortality alone. Our estimates for the total social valuation can be obtained by adding the social marginal willingness to pay for changes in adult and child mortality risks in Table 3. Using the Jenkins et al. (2001) numbers, the total aggregate value (in US\$ billions) is \$4891. For Mount et al. (2000) and Blomquist et al. (1996), the total value is \$5958 and \$7010, respectively. These values correspond to at least 1.7 times the U.S. aggregate GDP in 1965. From this, between \$1 trillion and \$3 trillion come directly from the effects of reductions in child mortality, through altruism and fertility, on the welfare of adults.

It is useful to contrast the estimates obtained in Table 3 with valuations based on alternative strategies commonly employed in the literature. Assuming that current children value themselves as if they were adults, and ignoring future generations, the estimated aggregate social value becomes \$5131 (in US\$ billions). The previous figures suggest that ignoring the value parents place on gains to be experienced by future generations gives something close to a lower bound to the social value of changes in mortality. Alternatively, if we assume that an interest rate of 3% also serves as an intergenerational discount rate, and that the size of future cohorts remains constant, we obtain an aggregate social value of \$8332. This estimate, representing the application of the strategy adopted by Murphy and Topel (2003) to our context, is larger than our highest total valuation, indicating that an intergenerational discount based on altruism and endogenous fertility exceeds the personal discount rate. In some sense, the traditional strategies adopted in the literature seem to provide lower and upper bounds to our estimate of the social value of mortality changes.

Finally, though the results are quantitatively large and point to the relevance of altruism and fertility in value of life-type calculations, we must repeat that several dimensions of our empirical implementation tend to underestimate the very factors that

¹⁸ The total value of changes in mortality risks for a young adult ranges from \$29,558 (using Jenkins et al., 2001) to \$56,720 (using Blomquist et al., 1996). Estimates based on Mount et al. (2000) generate a mid-range value of \$44,205. The traditional methodology applied in our context would generate a young adult's value grossly underestimated (\$19,218).

Table 3

Value of mortality reductions in the U.S. between 1965 and 1995, different sets of parameters

Parameterization	MWP_{18}^C (1996 \$)	MWP_{18}^A (1996 \$)	Social MWP^C (Bill., 1996 \$)	Social MWP^A (Bill., 1996 \$)
Blomquist et al. (1996)	36,449	20,272	2942	4,068
Mount et al. (2000)	23,670	20,536	1911	4,074
Jenkins et al. (2001)	11,697	17,861	944	3,947

Note: Authors' calculations based on parameter values presented in Table 2, benchmark variables from 1965 presented in Table 1, and the model discussed in the previous section.

we are trying to highlight. Indeed, we see the evidence as an extremely strong case in favor of the incorporation of fertility and altruism into the analysis.

4. Concluding remarks

This paper develops and applies an extended version of the “value of life” methodology that incorporates altruism and fertility decisions. In doing so, we are able to account for the value that parents attribute to reductions in child mortality and to incorporate the welfare of future generations in the evaluation of current health policies. We show that, under these circumstances, the value of adult mortality changes can be decomposed into three factors: the consumption factor from the traditional value of life specification; a fertility factor, which accounts for the welfare improvements related to the higher probability of having children; and an altruism factor, which accounts for the fact that mortality reductions will also be enjoyed by all future generations. The value of changes in child mortality, on its turn, depends on the effect of child mortality reductions on the final costs of child production, on the uncertainty regarding the number of surviving children, and on the emotional loss associated with the death of a child. As with adult mortality, child mortality reductions are also valuable because they are experienced by all future generations.

Applying our framework to value the changes in mortality observed in the U.S. between 1965 and 1995, we show that altruism and fertility can double the estimated welfare gain of a young adult. In relation to the aggregate social value, the increase is between 24% and 72%.

The paper also has implications for the evaluation of any other dimension of public policy that may have long-lasting effects (environmental policy, infrastructure development, technological improvements, etc.). The same principle used here to evaluate the welfare gain of future generations can also be applied in these other contexts. Altruism is a natural link between generations that immediately brings together the welfare gains of all future cohorts. And, from the perspective of the population currently alive, these gains are incorporated in a way that reflects precisely the value that public policy should attribute to future populations.

Finally, our application of state-dependent preferences to model the way parents value the utility of children and their survival finds support on anecdotal observations. When deciding over the custody of a baby, King Solomon—in the biblical example (1 Kings 3:16–28)—threatens to split the live baby in half. When one woman accepts the compromise, but the other claims that the baby does not belong to her (for the sake of her child's life), King Solomon realizes that the latter is the mother. Obviously, the loss of utility for the mother from a change between states (child death) is much larger than the loss, conditional on the child being alive, from a transfer of custody. Similarly, there are large differences between lawsuit awards given for wrongful deaths of children and awards given for wrongful births (due to contraceptive failures). Awards for wrongful births are low and mostly based on the costs of raising a child, while awards for wrongful deaths are typically extremely high, at values roughly near our monetary measure of emotional loss (see Zelizer, 1994). This pattern supports the idea that the utility cost of losing an existing child is indeed very different from the marginal cost of acquiring an additional child (which is directly related to the marginal benefit of an additional child), even when replacement through additional births is possible.

In our calibrated model, child survival dominates any concern about welfare of children conditional on survival. Therefore, our state-dependent formulation is entirely consistent with the views expressed in the examples above. Also from this perspective, the theoretical setup suggested here seems to be a promising benchmark for the incorporation of altruism and fertility—and therefore future generations—in the evaluation of public policies.

Appendix A. Calibration of α , ϕ and M_c

For the calibration of α , ϕ and M_c we employ information on marginal willingness to pay for changes in mortality risk. Calibrations based on the “value of statistical life” are standard in the literature (see, for example, Murphy and Topel, 2003) because inferring risk valuations from consumption patterns is problematic as risk taking cannot be separated from actual consumption decisions (see Rosen, 1981, 242–244) and Rosen (1988, 297).

Define the “value of a statistical life” as the willingness to pay for a one-shot unit reduction of the probability of death at some specific age (as Murphy and Topel, 2003), without any similar change for future generations. This definition is in line with the empirical interpretation of the “value of a statistical life” and also with the data typically used to estimate this number. Suppose that, for both adult and child mortalities, we are considering decisions when adults are about to realize their fertility choices (age τ ,

which is closest from the perspective of the model to adults that are raising children). From Eqs. (6) and (7), the “value of a statistical life” for adults and children are given by, respectively:

$$\text{VSL}_T^A = \left(\frac{1}{\varepsilon_c} - 1\right) c \int_T^\infty S(t, \tau) dt + (b + p_c e) \left(\frac{1}{\varepsilon_n} - 1\right) n_T, \text{ and}$$

$$\text{VSL}_T^C = \frac{1}{p_c} \left\{ b n_T + \frac{(1-\phi)}{2 + (1-\phi)^2 \mu^2} (b + p_c e) + \frac{\alpha p_c M_c}{u_c(c(t))} n_T \right\}.$$

With estimates of the “value of a statistical life” for adults and children, the expressions for VSL_T^A and VSL_T^C above define a system of two non-linear equations on α , ϕ , and M_c . First order conditions for c_t and n_T define a third equation, given by

$$\alpha \left\{ \frac{\partial \mathbb{E} [N_T^\phi]}{\partial n_T} V_{T+1} - M_c (1-p_c) \right\} = (b + e p_c) u_c(c_t).$$

In a stationary environment, V_{T+1} can be substituted by the closed form solution for the constant steady-state value function. Together with the two equations for VSL_T^A and VSL_T^C , this expression defines a system of three non-linear equations on three unknowns (α , ϕ and M_c).

Blomquist et al. (1996), Mount et al. (2000), and Jenkins et al. (2001) present estimates for VSL_T^A and VSL_T^C for the mid-1990's (see Table 1). The other variables used in the calibration are set to their 1995 values. Additional parameters, functional forms, and sources of variables used are discussed in the text. The solution to this problem for each different set of estimates of VSL_T^A and VSL_T^C generates the values of α , ϕ , and M_c presented in Table 1.

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